

Engineering Statistics
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Lecture 42
Hypothesis Testing, Bayes Test

So, we started talking about hypothesis testing. And we said that hypothesis is a statement about a population parameter.

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Hypothesis Testing

Definition: A Hypothesis is a statement about a population parameter

Definition: Two complementary hypothesis in a hypothesis testing are called *null hypothesis* and *alternate hypothesis*, denoted as H_0 and H_1 , respectively.

General form of hypothesis testing

$H_0 : \theta \in \Theta_0$ null hypothesis
 $H_1 : \theta \in \Theta_0^c$ null hypothesis

$H_0 : \theta \leq \theta_0$
 $H_1 : \theta > \theta_0$

for some threshold θ_0

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And we said in hypothesis there are actually two hypothesis, H_0 and H_1 , which are complementary. And the H_0 hypothesis is basically a statement like my parameter theta belongs to some set big theta 0. And that we are calling it as null hypothesis. And alternative hypothesis is my theta belongs to the complement class, which I am calling it as null hypothesis. And if my theta's are like real line, I may put a partition somewhere. Like if my theta space is like a real line, I may say okay, I am put this as my theta 0 and ask whether I this is below this or above this.

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Hypothesis Testing contd..

Null hypothesis

x_1

Hypothesis testing procedure or hypothesis test is rule that prescribes

1. For which sample values the **decision** is made to accept H_0 as true
2. For which sample values H_0 is rejected and H_1 is accepted as true

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So, in hypothesis testing, we are going to come up with a rule that tells whether my hypothesis need to be accepted or rejected. So, the rule is going to give me tell me how to make decisions.

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Methods of tests: Likelihood Ratio Test (LRT)

Likelihood ratio test statistic for testing $H_0 : \theta \in \Theta_0$ and $H_1 : \theta \in \Theta_0^c$ is

$$\lambda(\mathbf{x}) = \frac{\sup_{\theta \in \Theta_0} L(\theta|\mathbf{x})}{\sup_{\theta \in \{\Theta_0, \Theta_0^c\}} L(\theta|\mathbf{x})}$$

A Likelihood Ratio Test (LRT) is any any test that has the reject region of the form $\{x : \lambda(\mathbf{x}) \leq c\}$ for some $c \in (0, 1)$.

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And one of the possible decisions criteria we came up, is based on likelihood ratio test, which we call LRT. So, LRT is defined as a ratio of 2 quantities, which are defined in terms of your likelihood function. The numerator is optimizing your likelihood function over theta naught, which is basically the parameters belonging to your null hypothesis. And that denominator is optimizing over the entire parameter space, which is theta 0 as well as theta 0 complement.

So, this ratio whatever we are going to get for a given samples x be denoted as L of x sorry lambda of x .

And our decision is going to be now if lambda x is going to be less than or equals to c , I am going to reject it to be null hypothesis and accept it as an alternate hypothesis. And notice that there is a c here which is the parameter that you are going to set for making this decision and what will be the range of lambda x here, obviously, denominator is going to be larger than the numerator my likelihood functions are positive. So, this ratio is going to be between 0 1. So, my c will be also for some range between c 0 1 we will take.

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Example: $x = (x_1, x_2, \dots, x_n)$ $x_i \sim \text{Ber}(p)$

$$L(p|x) = \frac{p^{\sum x_i} (1-p)^{n-\sum x_i}}{\binom{n}{\sum x_i}}$$

$$\log(L(p|x)) = \sum x_i \log p + (n - \sum x_i) \log(1-p)$$

✓ $H_0: p = 1/2 = \theta_0$

✓ $H_1: p \neq 1/2 = \theta_0^c = \{p \in [0,1], p \neq 1/2\}$

$$\lambda(x) = \frac{\sup_{p \in \theta_0^c} L(x|p)}{\sup_{p \in [0,1]} L(x|p)} = \frac{\binom{n}{\sum x_i} (1/2)^n}{\binom{n}{\sum x_i} \left(\frac{\sum x_i}{n}\right)^{\sum x_i} \left(1 - \frac{\sum x_i}{n}\right)^{n - \sum x_i}} = \frac{(1/2)^n}{\left(\frac{\sum x_i}{n}\right)^{\sum x_i} \left(1 - \frac{\sum x_i}{n}\right)^{n - \sum x_i}}$$

$$\lambda(x) = c$$

So we calculated last time this example about the binomial random variable. Sorry, we did it for a Bernoulli random variable with parameter p and our hypothesis were whether it is a fair coin or an unfair coin. So, what we will do is, today we will do one more exercise. On the...

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Example 2 $X = (x_1, x_2, \dots, x_n)$ $x_i \sim N(\mu, \sigma^2)$ (known σ^2)

$H_0: \theta \leq \theta_0$
 $H_1: \theta > \theta_0$

$L(\theta|x) = \frac{1}{(2\pi\sigma^2)^{n/2}} \exp\left\{-\frac{\sum (x_i - \theta)^2}{2\sigma^2}\right\}$

$\lambda(x) = \frac{\sup_{\theta \leq \theta_0} L(\theta|x)}{\sup_{\theta} L(\theta|x)}$

diff expand: $-\frac{2 \sum (x_i - \theta)(-1)}{2\sigma^2} = \frac{\sum x_i - n\theta}{\sigma^2}$

Example 2 $X = (x_1, x_2, \dots, x_n)$ $x_i \sim N(\mu, \sigma^2)$ (unknown σ^2)

$H_0: \theta = \theta_0$
 $H_1: \theta \neq \theta_0$

$L(\theta|x) = \frac{1}{(2\pi\sigma^2)^{n/2}} \exp\left\{-\frac{\sum (x_i - \theta)^2}{2\sigma^2}\right\}$

$\lambda(x) = \frac{\sup_{\theta = \theta_0} L(\theta|x)}{\sup_{\theta} L(\theta|x)} = \frac{\exp\left\{-\frac{\sum (x_i - \theta_0)^2}{2\sigma^2}\right\}}{\exp\left\{-\frac{\sum (x_i - \bar{x})^2}{2\sigma^2}\right\}} = \exp\left\{-\frac{\sum (x_i - \theta_0)^2 - \sum (x_i - \bar{x})^2}{2\sigma^2}\right\} \leq c$

Given $c \in (0, 1)$

$R = \{x: \lambda(x) \leq c\}$

$-\frac{\sum (x_i - \theta_0)^2 + \sum (x_i - \bar{x})^2}{2\sigma^2} \leq (\log c)$

$\sum (x_i - \theta_0)^2 - \sum (x_i - \bar{x})^2 \geq -2\sigma^2 \log c$

$\sum x_i^2 + n\theta_0^2 - 2\sum x_i \theta_0 - [\sum x_i^2 + n\bar{x}^2 - 2\sum x_i \bar{x}] \geq -2\sigma^2 \log c$

$n(\theta_0 - \bar{x})^2 - 2n\bar{x}(\theta_0 - \bar{x}) = n(\theta_0 - \bar{x})(\theta_0 - \bar{x} - 2\bar{x})$

Let us say your samples now, let us take and sample and assume that this x_i 's are coming from μ sigma square. And you want to ensure that, so, okay fine, your x_i 's are IID coming from gaussian distribution with parameter μ and sigma square and now your goal is to identify your hypothesis on this parameter μ and let us assume that the sigma square is known. And now I want to make two hypothesis, my null hypothesis is θ is going to be less than or equals to some θ_0 and my H_1 is θ is going to be greater than this θ_0 , fine.

Now, first I want to know find out a likelihood ratio apply a likelihood ratio test on this and I want to find out that (likeli) what is that likelihood ratio and for that I need to first find out

my likelihood function. So, what is the likelihood function of the θ given x , we know this is nothing but $\frac{1}{(2\pi\sigma^2)^{n/2}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \theta)^2\right)$. So, now let us write $L(\theta|x)$ this is $\sup_{\theta} L(\theta|x)$ divided, θ less than or equals to θ not, this whole quantity and \sup over entire θ $L(\theta|x)$.

Now first let us do the bottom one, denominator one because that is unconstrained whereas the numerator is the constraint one. So, if you try to optimize our θ , what is the θ that will optimize this likelihood function? Sample mean. So, that is the one, the one that denominator is going to be a sample mean. What is the one, that is going to optimize the numerator? Why? Anybody can think why that is going to optimize a θ not? So, just think about this, whether this quantity here is increasing or decreasing and θ , decreasing? Let us say if you increase θ $x_i - \theta$ is going to decrease and with a minus sign increase.

So what is the effect of increasing θ on this quantity? Is it in increase? But notice that it is a square here it is not a linear term here. Does the square bother? Why do not I just work out quickly? What is that? Say this is a , the only parameter θ is coming here. I do not need to worry about this guy is a constant for me. Because σ^2 is known. And now only thing you have to optimize is this quantity.

And, here, I am interested in maximizing but if you ignore this minus sign, you have to just minimize this quantity. Find out what is that θ that minimizes this, maybe let us quickly do that. So if I differentiate this, this is going to be simply $x_i - \theta$ times minus 1. I am just differentiating this quantity here everybody is in sync with me. Differentiating exponent I am just trying to differentiate the exponent, this is the exponent here or maybe let me keep that minus here. This minus I also kept, so $2\sigma^2$.

Now, so this quantity becomes what is become minus minus summation $x_i - \theta$ not, sorry $n\theta$, now, you have also the constant θ is less than or equal to θ not, how we are going to handle this constraint or 2 also get knocked out fine. So, how you do? How do you handle this θ less than or equal to θ constraint, you want to construct a language here, or not necessary?

Let us simplify this problem, let us say instead of this I will make your life easier, let us only consider this problem here θ is equals to θ naught and θ is not equals to θ not, is this simpler? I am only considering the hypothesis whether θ is θ not or not. So,

then in the numerator there is nothing to optimize is actually a theta naught. And then what I will do maybe I will because of this simplification I will not go into all this circus.

So, again the numerator is going to be simply I will get rid of this constant because that anyway get counts canceled in the numerator and denominator. So, exponential minus xi minus theta not square 2 sigma square divided by exponential minus xi and what is that theta not that is maximizing? Sorry what is that theta that is maximizing this, that is exactly \bar{x} , everybody agree with this? Now, if you simplify this is nothing but exponential minus xi minus theta not square minus xi minus \bar{x} whole squared divided by 2 sigma square.

This is my lambda, this is my likelihood ratio test. Now, I want to set my rejection region as or x such that lambda of x is less than or equals to c for some given c , I am going to set given c which is belonging to $[0, 1]$, maybe I should keep it open interval $(0, 1)$. So, now, let us write this what is going to happen?

So, now, if I want to now equate this to less than or equals to c , if I do this what is the condition I am going to get minus xi minus theta whole square plus the summation will retain xi minus \bar{x} whole square and I will take the exponential other side this is going to be $\log c$, and there is also that 2 sigma square factor is there, but I think I should not take it like this I will just keep it like this, is this side inequality fine.

Now, what I will do is I will take this minus sign with 2 sigma square on the other side. So, you will end up xi minus why this is supposed to be theta not, this is a fixed theta not, theta not minus xi minus \bar{x} whole square now, this inequality become minus greater than or equal to, minus 2 sigma square log. Now, is there any simpler expression for this?

Student: Audio not clear.

Professor: Say, that again xi...

Student: \bar{x} minus theta.

Professor: \bar{x} minus theta not whole square...

Student: Audio not clear.

Professor: Oh, that is it, this will just happen to be \bar{x} minus theta not. Do you agree that it is going to be simply \bar{x} minus theta not this left-hand side?

Students: Audio not clear.

Professor: Why do not you quickly work out? I also want to work out this. Let us do, so, the left-hand side is going to be $\sum \xi^2 + n\theta_0^2 - 2\sum \xi\theta_0$, the first term. And the second term is going to be $\sum \xi^2 + n\bar{x}^2 - 2\bar{x}\sum \xi$, this whole thing and now this guy gets knocked up with this and I will club these two guys together, $n\theta_0^2 - n\bar{x}^2$.

And now I am going to these two guys, so these two guys how I can club. So, 2, I am going to divide this guy by n this sum because θ_0 is a constant, I will divide it by n . So I am going to get $\bar{x}\theta_0$ naught this guy to this and this guy, here \bar{x} is also constant for me, from the data. So, that is going to be $-2\bar{x}\sum \xi$, you are right, this should be plus. And this you simplify this is going to be $n\theta_0^2 - n\bar{x}^2 - 2n\bar{x}\sum \xi$, $2n\bar{x}\sum \xi$ I have taken and $\theta_0 - \bar{x}$, is this correct? That is it, now we have this it is not as simple as, as you said that it is simply $\theta_0 - \bar{x}$...

Students: Audio not clear.

Professor: You want to write it as $\theta_0 - \bar{x}$ into $\theta_0 + \bar{x}$, I mean product of $\theta_0 - \bar{x}$ into the $\theta_0 + \bar{x}$, let us do that. So, I am going to take \bar{x} outside and n is anyway common this is going to be $\theta_0 + \bar{x}$ and minus, n is already out to \bar{x} . So, fine this is going to be $n(\theta_0 - \bar{x})(\theta_0 + \bar{x})$, is it coming to that guy, $\theta_0 - \bar{x}$ is that correct? So, you are his friend want to prove him right, fine, then this is going to be $n(\theta_0 - \bar{x})^2$ whole square,

And this, so, final condition is $\theta_0 - \bar{x}$ whole this should be greater than or equals to $\sqrt{-2\log c/n}$, this is my condition. If you want to further simplify, you can take the square root both sides divide take this n on the other side and you will get everything just in terms of if \bar{x} is greater than something, but this is the condition you are going to get, if your \bar{x} is such that this condition happens.

Then you are going to reject it that, that is equals θ_0 equals to θ_0 . And if this condition is not true, then you are going to accept that that is indeed has a parameter θ_0 equals to θ_0 , fine. Any questions on this, now, let us look into.

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Sufficient Statistics and LRT

Let $T(\mathbf{X})$ be a sufficient statistic of sample $\mathbf{X} = \{X_1, X_2, \dots, X_n\}$ drawn from population $f(\mathbf{x}|\theta)$

$$\lambda(\mathbf{x}) = \frac{\sup_{\theta \in \Theta_0} L(\theta|\mathbf{x})}{\sup_{\theta \in \Theta} L(\theta|\mathbf{x})}$$

$$= \frac{\sup_{\theta \in \Theta_0} f(\mathbf{x}|\theta)}{\sup_{\theta \in \Theta} f(\mathbf{x}|\theta)}$$

$$= \frac{\sup_{\theta \in \Theta_0} h(\mathbf{x})g(T(\mathbf{x})|\theta)}{\sup_{\theta \in \Theta} h(\mathbf{x})g(T(\mathbf{x})|\theta)}$$

$$= \frac{\sup_{\theta \in \Theta_0} g(T(\mathbf{x})|\theta)}{\sup_{\theta \in \Theta} g(T(\mathbf{x})|\theta)}$$

$$= \frac{\sup_{\theta \in \Theta_0} L^*(\theta|T(\mathbf{x}))}{\sup_{\theta \in \Theta} L^*(\theta|T(\mathbf{x}))} = \lambda(T(\mathbf{x}))$$

$g(T(\mathbf{x}))$
 $\tau' = g \circ T(\mathbf{x})$
 $R_c = \{\mathbf{x} : \lambda(\mathbf{x}) \leq c\}$

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If I want to start with the likelihood ratio, sorry if I have if I already have a sufficient statistic and if I want to use it, what is the connection between that and my likelihood ratio test? So, let us say T of \mathbf{x} is a sufficient statistics on my sample \mathbf{x} , which is drawn from a population $f(\mathbf{x}|\theta)$ given θ . Now, our likelihood ratio test is this numerator and denominator. And by definition, this L of θ \mathbf{x} we are simply taking it to be the PDF. But now, we are talking about having a sufficient statistics, L sorry T , and we if you recall one of the properties of sufficient statistics is that we can split my PDF into two parts.

h and g , where h is a function of \mathbf{x} alone, and a g is a function of \mathbf{x} only through T . So I can write this in that case, I can quickly knock off this h of \mathbf{x} because it has nothing to do with my parameter θ . Now, what I have is g of T of \mathbf{x} , if d is a, T is a sufficient statistic, and g is another let us assume let us assume hypothetically g is a one to one function, what is this case g of T is, so, we have said like if T is a sufficient statistics and we are going to apply another function on this, this g composition T of \mathbf{x} , this is let us call, this is T' , this T' is again a sufficient statistic.

So, now we have a ratio in terms of two sufficient statistics, you can calculate I am simply writing it as L^* now, I have written it as L^* . Maybe I should have L^* now, this is a function of this given sufficient statistic. And now I am going to call got a new expression for my ratio test in terms of the sufficient statistics. That is what I am calling it as L of T of \mathbf{X} . So, I am this is just I am denoting, like when I am involving sufficient statistics T in my computation, this likelihood ratio is coming from this sufficient statistic T .

Now you can do all the things you want, like you, you just define your boundary rejection ratio, always like maybe we should write the rejection ratio with a subscript c because that depends on that parameter c, if you change your this is all x all x such that lambda x is going to be less than or equals to c, if you are going to change your c, your rejection region can potentially change.

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Bayes Tests:

In Baye's method, we have prior probability $P(\theta)$ and posterior probability $\pi(\theta|x)$ after observing sample x .

▶ $\pi(\theta|x)$ can be used to find $R_c = \{x: \lambda(x) \leq c\}$

→ $\Pr(\theta \in \Theta_0|x)$ and $\Pr(\theta \in \Theta_0^c|x)$

▶ We can assign

$$\Pr(\theta \in \Theta_0|x) = \Pr(H_0 \text{ is true } |x)$$

$$\Pr(\theta \in \Theta_0^c|x) = \Pr(H_1 \text{ is true } |x)$$

▶ Decision criteria

H_0 is true if $\Pr(\theta \in \Theta_0|x) \geq \Pr(\theta \in \Theta_0^c|x)$

H_1 is true otherwise

Rejection set $\{x: \Pr(\theta \in \Theta_0^c|x) > 1/2\}$.

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Sufficient Statistics and LRT

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$$\lambda(x) = \frac{\sup_{\theta \in \Theta_0} L(\theta|x)}{\sup_{\theta \in \Theta} L(\theta|x)}$$

$$= \frac{\sup_{\theta \in \Theta_0} f(x|\theta)}{\sup_{\theta \in \Theta} f(x|\theta)}$$

$$= \frac{\sup_{\theta \in \Theta_0} h(x)g(T(x)|\theta)}{\sup_{\theta \in \Theta} h(x)g(T(x)|\theta)}$$

$$= \frac{\sup_{\theta \in \Theta_0} g(T(x)|\theta)}{\sup_{\theta \in \Theta} g(T(x)|\theta)}$$

$$= \frac{\sup_{\theta \in \Theta_0} L^*(\theta|T(x))}{\sup_{\theta \in \Theta} L^*(\theta|T(x))} = \underline{\underline{\lambda(T(x))}}$$

$g(T(x))$
 $\tau' = g \circ \tau(x)$

$R_c = \{x: \lambda(x) \leq c\}$

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This is one way, now, this is kind of assuming a frequentist approach like assuming that. Theta is something some fixed parameter, which I do not know and that is the parameter which is governing my PDF. But we have always seen that instead of assuming theta to be fixed, we could be assuming that that is coming from another distribution itself. And in that

case, we use this Bayes approach. So the Bayes approach could be also used in the hypothesis testing. How to use it, let us see.

So, in the Bayes hypothesis testing what we do, we are going to assume some prior probability on my parameter θ , that is always we do and after like you observed some sample x maybe your posterior is going to be $\pi(\theta \text{ given } x)$. So that we have computed, how to compute the posterior given my prior information and an observation x . Now, I can now define my new hypothesis, instead of exactly whether it is except it belongs to this or this, instead of asking this binary decisions, we can make it probabilistic decision now, weather tell me the probability that the observation I am making is coming from this parameter space.

Earlier it was yes no, but now we are putting some probability on the yes and some probability on the no. So that is why we are saying now, once they have this posterior probability after observing data or your sample, you can ask the question, what is the probability that θ belongs to my null hypothesis given my observation x , is can you compute this probability? From your posterior probability? So once you know given x , this posterior probability is giving you probability distribution of this θ .

Now all I am asking is that it belongs to the set θ_0 . I should be able to compute this by integrating this probability or my space of θ_0 . And similarly, I should be also able to answer the question what is the probability that θ belongs to my complementarities, complementary set using my posterior probability. Now, we can assign instead of asking whether H_0 is true or not I can go in to ask the question, what is the probability that H_0 is true given x ? What is the probability that H_0 is false here. Sorry, this is correct.

This is H_0 and the what is the probability that H_1 is true given x . Now, these are probabilistic arguments. Now, the decision criteria has to come earlier my decision criteria was I was putting some threshold or my likelihood, sorry LIT function log, my earlier my decision criteria was like this, where I could probably compute my $\lambda(x)$, but here I am computing the probabilities given an observation x , what is the probability that H_0 is true and H_1 is true.

Now, what would be what would be a natural candidate to make a decision here whether H not is true or H_1 is true? You just may want to go like if this probability happens to be higher than this, then you will say that, H_0 is true, otherwise, H_1 is true. But then that so, you will

simply say that H_0 is true if θ is belonging to θ_0 given x is going to be larger than θ belonging to its complement set given x . And otherwise H_1 is true.

Alternatively, you can just say that, if I have to put it in this structure, I will compute this probability. And if this probability is greater than half, then I am going to reject it or basically accept the H_1 hypothesis. So, now, this λ is being replaced by the probability that my parameters are coming from a null hypothesis given my observation x . So, in the frequencies approach, we will just construct my likelihood ratio test.

And in this Bayesian approach, I am just going to compute this probability based on my posterior distributions. So this part on the Bayesian test, any questions on this? You can make it strictly greater than, if you are going to set it like this. Maybe you can just make it strictly greater than, that does not matter. Usually we were going to be like, that matters if your θ space is discrete, but most other times we will be working with θ space which is continuous in that case, it does not matter.