

Statistical Signal Processing
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Lecture 32
Recursive Least Squares (RLS) Adaptive Filter

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Let us recall

- In linear LS estimation, the observed data are represented as $X = A\theta + e$
where A is assumed to be a known full-rank matrix, $\theta = [\theta_1, \dots, \theta_K]^T$ is the parameter vector and e is the zero-mean random vector. The components of e are uncorrelated and have the same variance
- LS estimation minimizes the SSE given by
$$J(\theta) = (X - A\theta)^T (X - A\theta)$$
- Minimizing the SSE results in the normal equation
$$A^T A \hat{\theta}_{LS} = A^T X$$

Hello students. Welcome to this lecture on Recursive Least Squares, RLS Adaptive Filter. Let us recall in linear LS estimation the observed data are represented as X is equal to $A\theta$ plus e vector, where X is the data vector A is assumed to be known full rank matrix and θ is the parameter vector. So θ is the linear combination of parameter vector plus some noise vector and e is the zero-mean random vector.

The components of e are uncorrelated and have the same variance. So e is a random vector and its component has the same variance and zero mean and because of this random noise, this data is random. LS estimation minimizes this sum square error given by this cost function that is $J(\theta) = (X - A\theta)^T (X - A\theta)$ and minimizing this sum square error cost function, we get the normal equation that is given by $A^T A \hat{\theta}_{LS} = A^T X$.

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Let us recall...

The LS estimator is given by

$$\hat{\theta}_{LS} = (A^T A)^{-1} A^T X$$

where A^+ is the pseudo inverse of A .

Matrix inversion makes LS estimation computationally complex.

$\hat{\theta}_{LS}$ is the best linear unbiased estimator (BLUE)- it is unbiased and it has the lowest variance in the class of linear unbiased estimators.

This lecture will model the adaptive filtering problem as an LS estimation problem and present the *Recursive Least-Squares (RLS)* technique to estimate the filter parameters adaptively.

Now the LS estimator is given by this expression $\hat{\theta}_{LS}$ is equal to $A^T A$ inverse into $A^T X$ and this quantity $A^T A$ inverse into A^T is denoted by A^+ and is called the pseudo-inverse. A^+ is this pseudo-inverse. Therefore, $\hat{\theta}_{LS}$ is equal to pseudo inverse A into X . Now, we see that there is a matrix inversion. This matrix inversion makes LS estimation computationally complex.

We also note that $\hat{\theta}_{LS}$ is the best linear unbiased estimator BLUE. It is unbiased and it has the lowest variance in the class of linear unbiased estimators. So we will consider a class of linear unbiased estimators. In that class, $\hat{\theta}_{LS}$ has the lowest variance. This lecture will model the adaptive filtering problem as an LS estimation problem and present the recursive least square technique to estimate the filter parameters adaptively.

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Drawback of the LMS filters

❖ LMS algorithm is computationally simple, but has the following drawbacks:

- LMS convergence slow, particularly when the eigen values of the autocorrelation matrix R_y are widely spread. $\frac{\lambda_{max}}{\lambda_{min}} \gg 1$
 - Step size parameter is to be properly chosen
 - Excess mean-square error is high
- ❖ The RLS algorithm converges faster even when the eigen value spread of the autocorrelation matrix is large.
- ❖ The problem of selecting the step-size parameter is avoided.
- However, the RLS algorithm has higher computational complexity.

We will start with the drawback of the LMS filter. LMS algorithm is computationally very simple, but it has the following drawbacks. The major drawback is LMS convergence is very slow, particularly when the eigenvalues of the autocorrelation matrix R_y are widely spread, if λ_{max} divided by λ_{min} is much greater than 1. Then convergence is slow. Step size parameter μ is to be properly chosen.

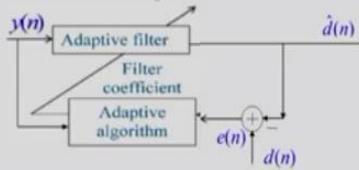
So that is one important factor, otherwise the algorithm may diverge. Excess mean square is high, particularly if step length is high, then this excess mean square error will be high. The RLS algorithm converges faster even when the eigenvalue spread of the autocorrelation matrix is large. So this is one important advantage of RLS algorithm. The problem of selecting the step size parameter is avoided in RLS algorithm. However, as we will see later, the RLS algorithm has higher computational complexity.

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LS approach

❖ Consider the FIR adaptive filter as shown in the figure. The LMS algorithm minimizes the instantaneous square error $e^2(n)$

$$\text{where } e(n) = x(n) - \mathbf{h}'(n)y(n) = x(n) - y'(n)\mathbf{h}(n)$$



❖ The LS approach minimizes the sum-square error between the desired signal $d(n)$ and the filtered output $\hat{d}(n)$.

❖ It considers all the available data up to the current time for determining the filter parameters.

❖ The filter is optimum with respect to all the available data in the least-squares sense. The associated LS estimation problem is solved recursively.

Let us see the FIR filter structure for adaptive filter implementation. So here y_n is the input signal. This is the FIR adaptive filter. Filter output is \hat{d}_n and it is compared with a desired signal d_n and $d_n - \hat{d}_n$ is the error. So that error is used to adaptively update the filter coefficient and those filter coefficients are used to filter this signal. We know that the LS approach minimizes the sum square error, SSE between the desired signal d_n and the filtered signal \hat{d}_n , because here error is between the desired signal and the filtered signal.

Therefore, LS approach will minimize this sum square of this error. It considers all the available data up to the current time for determining the filter parameters. This is the advantage of RLS algorithm that it considers all the available data up to the current time for determining the filter parameters. The filter is optimum with respect to all the available data in the least-squares sense. The associated LS estimation problem is solved recursively. Now this filter will be optimum with respect to all the previous data. That is one important aspect and this estimation problem is solved recursively.

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Weighted sum-square error

❖ Suppose $\mathbf{h}(n) = \begin{bmatrix} h_0(n) \\ h_1(n) \\ \vdots \\ h_{M-1}(n) \end{bmatrix}$ is the filter coefficient vector at instant n .

❖ The cost function of the RLS adaptive filter is the weighted SSE given by

$$\begin{aligned} \varepsilon(n) &= \sum_{k=0}^n \lambda^{n-k} e^2(k) \\ &= \sum_{k=0}^n \lambda^{n-k} (d(k) - \mathbf{y}'(k)\mathbf{h}(n))^2 \\ &= \sum_{k=0}^n \lambda^{n-k} (d(k) - \mathbf{h}'(n)\mathbf{y}(k))^2 \end{aligned}$$

where $e(k)$ is the filtering error at instant k due to $\mathbf{h}(n)$.

$0 < \lambda \leq 1$ is the weighing factor known as the forgetting factor.

Suppose $\mathbf{h}(n)$ that is equal to this filter coefficient vector given by $h_0(n)$ $h_1(n)$ upto $h_{M-1}(n)$. The cost function of RLS adaptive filter is the weighted sum square error. So here the error scores will be weighted. This is given by this; that is epsilon error is the sum squared error. This is equal to summation lambda to the power $n - k$ into $e^2(k)$, k going from 0 to n , so up to instant n from k is equal to 0 up to instant n , all the error squares are weighted and summed up.

And this $e(k)$, we can write as $d(k) - \mathbf{y}'(k)\mathbf{h}(n)$ vector transpose into $\mathbf{h}(n)$ vector and which can be also written as $d(k) - \mathbf{h}'(n)\mathbf{y}(k)$ vector transpose into $\mathbf{y}(k)$ vector. Now here $e(k)$ is the filtering error at instant k due to $\mathbf{h}(n)$. So here you see that filter at instant n , that is $\mathbf{h}(n)$ that filter is used to filter all the previous data $\mathbf{y}(k)$, k from 0 to n . So that way $e(k)$ is the filtering error at instant k due to filter $\mathbf{h}(n)$. This we have to note and this quantity lambda which lies between 0 and 1 is the weighing factor known as the forgetting factor. So this factor is known as the forgetting factor.

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Importance of λ

- Past filtering errors $e(k)$, $k = 0, 1, \dots, n$ are computed with the most recent filter $\mathbf{h}(n)$.
- Recent $e(n)$ is given more weightage. Past errors are given smaller and smaller weights. This allows the filter coefficients to adapt to the non-stationarity in the data.
- For stationary case $\lambda = 1$ can be taken
- $\lambda \cong 0.99$ is effective in tracking local non-stationarity.

Now let us see the importance of the forgetting factor lambda, no depth past filtering errors e_k , k going from 0 to n are computed with the most recent filter \mathbf{h}_n . This is one important observation we have already made. Recent error is given more weightage; past errors are given smaller and smaller weights. How this is so? Because this is lambda to the power $n - k$, when k is suppose equal to n , then lambda will be equal to 1.

And when k is 0 suppose, then it will be lambda to the power n and this will be a very small quantity, because lambda is a number less than. That way recent error is given more weightage; past errors are given smaller and smaller weights. This allows the filter coefficients to adapt to the non-stationarity in the data. If data is non-stationary, then the recent error will reflect that non-stationarity and since it is given more weightage, therefore we will get better estimation. For stationary case, lambda = 1 can be taken. Lambda approximately equal to 0.99 is effective in tracking local non-stationarity. So it should be less than 1 and we can take something around 0.99.

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Normal equation

❖ The LS estimation problem is

$$\text{Minimize } \varepsilon(n) = \sum_{k=0}^n \lambda^{n-k} (d(k) - \mathbf{y}'(k) \mathbf{h}(n))^2 \text{ with respect to } \mathbf{h}(n)$$

❖ The minimum is given by $\frac{\partial \varepsilon(n)}{\partial \mathbf{h}(n)} = \mathbf{0}$

$$\Rightarrow 2 \sum_{k=0}^n \lambda^{n-k} (d(k) - \mathbf{y}'(k) \mathbf{h}(n)) \mathbf{y}(k) = 0$$

$$\Rightarrow 2 \sum_{k=0}^n \lambda^{n-k} (d(k) \mathbf{y}(k) - \mathbf{y}(k) \mathbf{y}'(k) \mathbf{h}(n)) = 0$$

$$\therefore \left(\sum_{k=0}^n \lambda^{n-k} \mathbf{y}(k) \mathbf{y}'(k) \right) \mathbf{h}(n) = \sum_{k=0}^n \lambda^{n-k} d(k) \mathbf{y}(k)$$

which is the *normal equation* for the LS estimation.

Now the LS estimation problem is given by minimize; this is sum square error up to instant n. This is given by summation lambda to the power n - k into dk - yk transpose into hn whole square, k going from 0 to n. This is the sum square error. This sum square error we have to minimize with respect to hn vector. So each component of hn we have to use for minimization. Now the minimum is given by this condition del en del hn vector is equal to 0 vector.

So here differentiation with respect to hn vector and we can carry out the differentiation with respect to hn vector. We will get because of this square 2 will come and then this part will be; they are as dk - yk transpose into hn and then because this quantity also has term hn. If we differentiate with respect to hn, we will get yk here. Therefore, the minimum condition will be now given by 2 times summation lambda to the power n - k into dk - yk vector transpose into hn vector into yk, k going from 0 to n, that must be equal to 0.

Now this yk, we can take inside and we note that this quantity is a scalar. So yk we can bring here. Similarly, this quantity is also a scalar; yk we can bring here. This yk is a vector. Therefore, we will get, 2 will be there. Anyway 2 can be cancelled. Summation y to the power n - k into dk into yk vector - yk vector into yk vector transpose into hn, k going from 0 to n is equal to 0.

So now bringing this part to the right hand side, we will get summation lambda to the power n - k into yk into yk transpose, k going from 0 to n into hn is equal to summation lambda to the

power $n - k$ into dk that is a scalar into y_k vector, k going from 0 to n . So we have a matrix equation now and this equation is known as the normal equation for the LS estimation. Now this quantity, this is similar to the estimators for autocorrelation matrix.

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Normal equation ...

- ❖ Let us define $\hat{\mathbf{R}}_y(n) = \sum_{k=0}^n \lambda^{n-k} \mathbf{y}(k) \mathbf{y}'(k)$, which is an estimator for the autocorrelation matrix \mathbf{R}_y .
- ❖ Similarly $\hat{\mathbf{r}}_{dy}(n) = \sum_{k=0}^n \lambda^{n-k} d(k) \mathbf{y}(k)$ is an estimator for the cross-correlation vector $\mathbf{r}_{dy}(n)$
- ❖ Hence the normal equation can be rewritten as

$$\hat{\mathbf{R}}_y(n) \mathbf{h}(n) = \hat{\mathbf{r}}_{dy}(n)$$
- ❖ The solution to the normal equation is given by

$$\mathbf{h}(n) = \hat{\mathbf{R}}_y^{-1}(n) \hat{\mathbf{r}}_{dy}(n)$$
- ❖ Matrix inversion is involved, which makes the direct solution computationally complex. We look forward for a recursive solution.

So we will denote $\hat{\mathbf{R}}_y(n)$ equal to summation λ to the power $n - k$ into y_k into y_k transpose, k going from 0 to n . So this is a matrix. This is the estimator for the autocorrelation matrix \mathbf{R}_y . Similarly, we can define this quantity summation k going from 0 to n λ to the power $n - k$ into dk into y_k vector. So that way, this will measure the correlation between dk and component of y_k vector and then you are summing up from k is equal to 0 to n .

So that way, this is also an estimator for the cross correlation vector \mathbf{R}_{dy} . So we can now write $\hat{\mathbf{R}}_y(n)$ for this expression and $\hat{\mathbf{r}}_{dy}(n)$ for this vector. So hence the normal equation can be rewritten as $\hat{\mathbf{R}}_y(n) \mathbf{h}(n) = \hat{\mathbf{r}}_{dy}(n)$ vector. This matrix is invertible. Therefore, the solution of the normal equation is given by $\mathbf{h}(n) = \hat{\mathbf{R}}_y^{-1}(n) \hat{\mathbf{r}}_{dy}(n)$ vector. So that way here, the filter coefficients are to be obtained by matrix inversion.

We have to invert the estimated autocorrelation matrix. Here matrix inversion is involved, which makes the direct solution computationally complex, because finding inverse of a big matrix is computationally very complex. We look forward for a recursive solution.

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Recursive representation of $\hat{\mathbf{R}}_y(n)$

❖ $\hat{\mathbf{R}}_y(n)$ can be rewritten as follows

$$\begin{aligned}\hat{\mathbf{R}}_y(n) &= \sum_{k=0}^{n-1} \lambda^{n-k} \mathbf{y}(k) \mathbf{y}'(k) + \mathbf{y}(n) \mathbf{y}'(n) \\ &= \lambda \sum_{k=0}^{n-1} \lambda^{n-1-k} \mathbf{y}(k) \mathbf{y}'(k) + \mathbf{y}(n) \mathbf{y}'(n) \\ &= \lambda \hat{\mathbf{R}}_y(n-1) + \mathbf{y}(n) \mathbf{y}'(n) \\ \therefore \hat{\mathbf{R}}_y(n) &= \lambda \hat{\mathbf{R}}_y(n-1) + \mathbf{y}(n) \mathbf{y}'(n)\end{aligned}$$

$$\begin{aligned}\hat{\mathbf{R}}_y(n) &= \sum_{k=0}^n \lambda^{n-k} \mathbf{y}(k) \mathbf{y}'(k) \\ &= \sum_{k=0}^{n-1} \lambda^{n-k} \mathbf{y}(k) \mathbf{y}'(k) + \lambda^0 \mathbf{y}(n) \mathbf{y}'(n)\end{aligned}$$

❖ This shows that the autocorrelation matrix can be recursively computed from its previous values and the present data vector.

❖ Similarly, the cross-correlation vector is given by

$$\hat{\mathbf{r}}_{xy}(n) = \lambda \hat{\mathbf{r}}_{xy}(n-1) + d(n) \mathbf{y}(n)$$

First, we will see the recursive representation of $\hat{\mathbf{R}}_y(n)$. $\hat{\mathbf{R}}_y(n)$ can be rewritten as $\hat{\mathbf{R}}_y(n)$ is equal to λ times the summation of $\mathbf{y}(k) \mathbf{y}'(k)$, k going from 0 to $n-1$. Now, if we separate out the term involving $\mathbf{y}(n)$, so when we put k is equal to n , then that term will be $\lambda^0 \mathbf{y}(n) \mathbf{y}'(n)$. So if we take the first $n-1$ terms, then we will have summation λ times the summation of $\mathbf{y}(k) \mathbf{y}'(k)$, k going from 0 to $n-1$.

And last term when k is equal to n that will be given by $\lambda^0 \mathbf{y}(n) \mathbf{y}'(n)$ that is λ times the summation of $\mathbf{y}(k) \mathbf{y}'(k)$, k going from 0 to $n-1$ plus $\mathbf{y}(n) \mathbf{y}'(n)$. So, last term will be simply $\mathbf{y}(n) \mathbf{y}'(n)$. So therefore, $\hat{\mathbf{R}}_y(n)$ can be rewritten as this summation; summation λ times the summation of $\mathbf{y}(k) \mathbf{y}'(k)$, k going from 0 to $n-1$ plus $\mathbf{y}(n) \mathbf{y}'(n)$. Therefore, this term we can write as λ times the summation of $\mathbf{y}(k) \mathbf{y}'(k)$, k going from 0 to $n-1$ plus $\mathbf{y}(n) \mathbf{y}'(n)$.

Now this quantity is nothing but this is the estimator for the autocorrelation function at $n-1$, because it is from k is equal to 0 to $n-1$. Therefore, this part is $\hat{\mathbf{R}}_y(n-1)$. Therefore, we get λ times $\hat{\mathbf{R}}_y(n-1)$ plus this term, that is $\mathbf{y}(n) \mathbf{y}'(n)$. Therefore, we have established that the autocorrelation estimator at instant n is equal to λ times autocorrelation estimator at $n-1$ plus $\mathbf{y}(n) \mathbf{y}'(n)$.

So that way autocorrelation function can be incremented. This shows that the autocorrelation matrix can be recursively computed from previous values and the present data vector, because y_n is the present data vector. So y_n multiplied by y_n transpose, this part we are adding to the previous estimator. Similarly, the cross correlation vector is given by r_{dy} hat n . This can be recursively obtained λ times r_{dy} hat $n - 1 + d_n$ into y_n vector. So this way we can update this cross correlation function.

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Recursive solution of the normal equation

$$h(n) = [\hat{R}_Y(n)]^{-1} \hat{r}_d(n)$$

$$= (\lambda \hat{R}_Y(n-1) + y(n)y(n)')^{-1} \hat{r}_d(n)$$

❖ The sum matrix is obtained by adding a rank-one part at each instant.

How to get the inverse of the sum matrix?

❖ The **matrix inversion lemma** will be now useful.

Now we will see how we can get the recursive solution of the normal equation. We have $h(n)$ is equal to R_Y hat n inverse into r_{dy} hat n and this we can write as, because this term we can write recursively, so that is λ times R_Y hat $n - 1 + y_n$ into y_n transpose whole inverse into R_Y hat n . This sum matrix is obtained by adding a rank one part. This part is the rank one part, because this is only one column is there.

That column you are taking transpose and you are getting the matrix. This is a rank one matrix. So the sum matrix is obtained by adding a rank one part at each instant. Instant n , we are adding this rank one part. How to get the inverse of this sum matrix that is a problem. The matrix inversion lemma will be now useful. There is a powerful result what is known as the matrix inversion lemma. We will see that now.

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Matrix Inversion Lemma

❖ If A is an $M \times M$ non-singular matrix, \mathbf{u} and \mathbf{v} are two column vectors

$$(\mathbf{A} + \mathbf{uv}^T)^{-1} = \mathbf{A}^{-1} - \frac{\mathbf{A}^{-1} \mathbf{uv}^T \mathbf{A}^{-1}}{1 + \mathbf{v}^T \mathbf{A}^{-1} \mathbf{u}}$$

provided $1 + \mathbf{v}^T \mathbf{A}^{-1} \mathbf{u} \neq 0$

❖ This formula is known as the Sherman-Morrison formula. Once \mathbf{A}^{-1} is known, the inverse of the rank-one increment $(\mathbf{A} + \mathbf{uv}^T)^{-1}$ can be easily computed.

If A is an M by M non-singular matrix, so that inverse exists, \mathbf{u} and \mathbf{v} are two column vectors, then inverse of A plus \mathbf{uv}^T is equal to $A^{-1} - \frac{A^{-1} \mathbf{uv}^T A^{-1}}{1 + \mathbf{v}^T A^{-1} \mathbf{u}}$, provided this denominator term $1 + \mathbf{v}^T A^{-1} \mathbf{u}$ is not equal to 0. In this way, we can get the inverse of this sum in terms of the inverse of A and the matrix \mathbf{u} and \mathbf{v} itself.

So there is no other inversion, except A^{-1} and if we know A^{-1} , then this can be computed algebraically like this. This formula is known as the Sherman-Morrison formula. Once A^{-1} is known, the inversion of the rank one increment $A + \mathbf{uv}^T$ inverse can be easily computed. This is the advantage of the matrix inversion lemma. So suppose if we know A^{-1} , then inverse of the rank one increment, so $A + \mathbf{uv}^T$ that can be easily obtained by this formula.

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Computation of $\hat{\mathbf{R}}_Y^{-1}(n)$

• We have

$$(\mathbf{A} + \mathbf{u}\mathbf{v}')^{-1} = \mathbf{A}^{-1} - \frac{\mathbf{A}^{-1} \mathbf{u}\mathbf{v}' \mathbf{A}^{-1}}{1 + \mathbf{v}' \mathbf{A}^{-1} \mathbf{u}}$$

❖ Taking $\mathbf{A} = \lambda \hat{\mathbf{R}}_Y(n-1)$ and $\mathbf{u} = \mathbf{v} = \mathbf{y}(n)$, we get

we will have

$$\begin{aligned} \hat{\mathbf{R}}_Y^{-1}(n) &= (\lambda \hat{\mathbf{R}}_Y(n-1) + \mathbf{y}(n)\mathbf{y}(n)')^{-1} \\ &= \frac{1}{\lambda} \hat{\mathbf{R}}_Y^{-1}(n-1) - \frac{\frac{1}{\lambda} \hat{\mathbf{R}}_Y^{-1}(n-1) \mathbf{y}(n) \mathbf{y}'(n) \frac{1}{\lambda} \hat{\mathbf{R}}_Y^{-1}(n-1)}{1 + \mathbf{y}'(n) \frac{1}{\lambda} (\hat{\mathbf{R}}_Y^{-1}(n-1)) \mathbf{y}(n)} \\ &= \frac{1}{\lambda} \left(\hat{\mathbf{R}}_Y^{-1}(n-1) - \frac{\hat{\mathbf{R}}_Y^{-1}(n-1) \mathbf{y}(n) \mathbf{y}'(n) \hat{\mathbf{R}}_Y^{-1}(n-1)}{\lambda + \mathbf{y}'(n) \hat{\mathbf{R}}_Y^{-1}(n-1) \mathbf{y}(n)} \right) \end{aligned}$$

Now, let us see how to find out the inverse of the autocorrelation matrix. So we have this result. Matrix inversion lemma taking A is equal to suppose $\lambda \hat{\mathbf{R}}_Y(n-1)$ and $\mathbf{u} = \mathbf{v} = \mathbf{y}(n)$ vector. So what we will get then? The inverse of $\hat{\mathbf{R}}_Y(n)$ is given by the inverse of $\lambda \hat{\mathbf{R}}_Y(n-1) + \mathbf{y}(n)\mathbf{y}(n)'$ and now using these results A is equal to $\lambda \hat{\mathbf{R}}_Y(n-1)$ and $\mathbf{u} = \mathbf{v} = \mathbf{y}(n)$, we will get this inversion, will be equal to by using this formula $\frac{1}{\lambda} \hat{\mathbf{R}}_Y^{-1}(n-1)$ minus, now numerator this inverse of this quantity $\frac{1}{\lambda} \hat{\mathbf{R}}_Y^{-1}(n-1) \mathbf{y}(n) \mathbf{y}'(n) \frac{1}{\lambda} \hat{\mathbf{R}}_Y^{-1}(n-1)$ into $\mathbf{y}(n) \mathbf{y}'(n)$.

That is $\frac{1}{\lambda} \hat{\mathbf{R}}_Y^{-1}(n-1)$, then \mathbf{u} and \mathbf{v} are equal $\mathbf{y}(n)$ into $\mathbf{y}(n)'$ into again inverse that is $\frac{1}{\lambda} \hat{\mathbf{R}}_Y^{-1}(n-1)$ divided by now denominator term $1 + \mathbf{y}'(n) \frac{1}{\lambda} \hat{\mathbf{R}}_Y^{-1}(n-1) \mathbf{y}(n)$. We can take $\frac{1}{\lambda}$ outside. So we will get $\frac{1}{\lambda} \left(\hat{\mathbf{R}}_Y^{-1}(n-1) - \frac{\hat{\mathbf{R}}_Y^{-1}(n-1) \mathbf{y}(n) \mathbf{y}'(n) \hat{\mathbf{R}}_Y^{-1}(n-1)}{\lambda + \mathbf{y}'(n) \hat{\mathbf{R}}_Y^{-1}(n-1) \mathbf{y}(n)} \right)$.

Divided by denominator will be, because λ if we multiply by λ this one will be λ , so $\lambda + \mathbf{y}'(n) \hat{\mathbf{R}}_Y^{-1}(n-1) \mathbf{y}(n)$. So this is the expression for $\hat{\mathbf{R}}_Y^{-1}(n)$ and which is obtained from the inverse at instant $n-1$ and it is updated using the current data.

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Computation of $\hat{\mathbf{R}}_y^{-1}(n)$...

Denote $\mathbf{P}(n) = \hat{\mathbf{R}}_y^{-1}(n)$. Then

$$\begin{aligned}\mathbf{P}(n) &= \frac{1}{\lambda} \left(\mathbf{P}(n-1) - \frac{\mathbf{P}(n-1)\mathbf{y}(n)\mathbf{y}'(n)\mathbf{P}(n-1)}{\lambda + \mathbf{y}'(n)\mathbf{P}(n-1)\mathbf{y}(n)} \right) \\ &= \frac{1}{\lambda} (\mathbf{P}(n-1) - \mathbf{k}(n)\mathbf{y}'(n)\mathbf{P}(n-1)) \\ \therefore \mathbf{P}(n) &= \frac{1}{\lambda} (\mathbf{P}(n-1) - \mathbf{k}(n)\mathbf{y}'(n)\mathbf{P}(n-1))\end{aligned}$$

where $\mathbf{k}(n)$ is called the 'gain vector' and given by

$$\mathbf{k}(n) = \frac{\mathbf{P}(n-1)\mathbf{y}(n)}{\lambda + \mathbf{y}'(n)\mathbf{P}(n-1)\mathbf{y}(n)}$$

From the above, we get

$$\lambda\mathbf{k}(n) = \mathbf{k}(n)\mathbf{y}'(n)\mathbf{P}(n-1)\mathbf{y}(n) - \mathbf{P}(n-1)\mathbf{y}(n)$$

Now, we will denote this term $\hat{\mathbf{R}}_y^{-1}(n)$ that is inverse of the autocorrelation estimator as \mathbf{P}_n matrix. This is the standard notation \mathbf{P}_n is equal to the inverse of the estimated autocorrelation matrix. \mathbf{P}_n is equal to $\hat{\mathbf{R}}_y^{-1}$ and then from this, this \mathbf{P}_n will be equal to $\frac{1}{\lambda} \mathbf{P}_{n-1} - \frac{\mathbf{P}_{n-1} \mathbf{y}(n) \mathbf{y}'(n) \mathbf{P}_{n-1}}{\lambda + \mathbf{y}'(n) \mathbf{P}_{n-1} \mathbf{y}(n)}$. So that way, we write \mathbf{P}_n is equal to $\frac{1}{\lambda} \mathbf{P}_{n-1} - \frac{\mathbf{P}_{n-1} \mathbf{y}(n) \mathbf{y}'(n) \mathbf{P}_{n-1}}{\lambda + \mathbf{y}'(n) \mathbf{P}_{n-1} \mathbf{y}(n)}$.

Now, this part $\frac{\mathbf{P}_{n-1} \mathbf{y}(n) \mathbf{y}'(n) \mathbf{P}_{n-1}}{\lambda + \mathbf{y}'(n) \mathbf{P}_{n-1} \mathbf{y}(n)}$, this part we will denote by a vector that is gain vector \mathbf{k}_n . Therefore, we can let $\mathbf{P}_n = \frac{1}{\lambda} \mathbf{P}_{n-1} - \mathbf{k}_n \mathbf{y}'(n) \mathbf{P}_{n-1}$ that is the gain vector multiplied by $\mathbf{y}'(n) \mathbf{P}_{n-1}$. So this is the expression for recursive estimation of \mathbf{P}_n . $\mathbf{P}_n = \frac{1}{\lambda} \mathbf{P}_{n-1} - \mathbf{k}_n \mathbf{y}'(n) \mathbf{P}_{n-1}$. This \mathbf{k}_n is called the gain vector and given by $\mathbf{k}_n = \frac{\mathbf{P}_{n-1} \mathbf{y}(n)}{\lambda + \mathbf{y}'(n) \mathbf{P}_{n-1} \mathbf{y}(n)}$. This is the way we define the gain vector.

So this is the updating for the inverse of the autocorrelation matrix and this is the updating for the gain vector and from this result $\mathbf{k}_n = \frac{\mathbf{P}_{n-1} \mathbf{y}(n)}{\lambda + \mathbf{y}'(n) \mathbf{P}_{n-1} \mathbf{y}(n)}$, we can easily show that, because we can multiply λ into \mathbf{k}_n and that will be equal to $\mathbf{k}_n \mathbf{y}'(n) \mathbf{P}_{n-1} \mathbf{y}(n) - \mathbf{P}_{n-1} \mathbf{y}(n)$ minus this part $\mathbf{P}_{n-1} \mathbf{y}(n)$. This result is also useful for us.

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Gain vector

- ❖ $\mathbf{k}(n)$ is important to interpret adaptation. It is also related to the current data vector $\mathbf{y}(n)$ by

$$\mathbf{k}(n) = \mathbf{P}(n)\mathbf{y}(n)$$

- ❖ To establish the above relation consider

$$\mathbf{P}(n) = \frac{1}{\lambda} (\mathbf{P}(n-1) - \mathbf{k}(n)\mathbf{y}'(n)\mathbf{P}(n-1))$$

- ❖ Multiplying by λ and post-multiplying by $\mathbf{y}(n)$ and simplifying we get

$$\begin{aligned}\lambda\mathbf{P}(n)\mathbf{y}(n) &= (\mathbf{P}(n-1) - \mathbf{k}(n)\mathbf{y}'(n)\mathbf{P}(n-1))\mathbf{y}(n) \\ &= \mathbf{P}(n-1)\mathbf{y}(n) - \mathbf{k}(n)\mathbf{y}'(n)\mathbf{P}(n-1)\mathbf{y}(n) \\ &= \lambda\mathbf{k}(n) \\ \therefore \mathbf{k}(n) &= \mathbf{P}(n)\mathbf{y}(n)\end{aligned}$$

The gain vector \mathbf{k}_n is important to interpret adaptation. It is also related to the current data vector \mathbf{y}_n by this relationship $\mathbf{k}_n = \mathbf{P}_n \mathbf{y}_n$ where \mathbf{P}_n is the inversion of the autocorrelation matrix. This result we can establish. We start with the expression for \mathbf{P}_n . This is the expression for \mathbf{P}_n ; that we are considering \mathbf{P}_n is equal to $\frac{1}{\lambda} (\mathbf{P}_{n-1} - \mathbf{k}_n \mathbf{y}_n^T \mathbf{P}_{n-1})$ and multiplying by λ and post multiplying by \mathbf{y}_n , because this is a matrix.

We are post multiplying by \mathbf{y}_n and simplifying we get $\lambda \mathbf{P}_n \mathbf{y}_n$ is equal to $\mathbf{P}_{n-1} \mathbf{y}_n - \mathbf{k}_n \mathbf{y}_n^T \mathbf{P}_{n-1} \mathbf{y}_n$, because we are post multiplying by \mathbf{y}_n . So that way, this will be equal to $\mathbf{P}_{n-1} \mathbf{y}_n$. This is $\mathbf{P}_{n-1} \mathbf{y}_n$ minus $\mathbf{k}_n \mathbf{y}_n^T \mathbf{P}_{n-1} \mathbf{y}_n$ into \mathbf{y}_n . Now this result we know. We have shown that this result is equal to $\lambda \mathbf{k}_n$. So therefore, we write here $\lambda \mathbf{P}_n \mathbf{y}_n$ is equal to $\lambda \mathbf{k}_n$. Therefore, \mathbf{k}_n will be equal to $\mathbf{P}_n \mathbf{y}_n$. So this result we will be using.

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Filter updates

$$\begin{aligned}
 \mathbf{h}(n) &= \left(\hat{\mathbf{R}}_Y(n) \right)^{-1} \hat{\mathbf{r}}_{dY}(n) \\
 &= \mathbf{P}(n) \left(\lambda \hat{\mathbf{r}}_{dY}(n-1) + d(n) \mathbf{y}(n) \right) \\
 &= \lambda \mathbf{P}(n) \hat{\mathbf{r}}_{dY}(n-1) + d(n) \mathbf{P}(n) \mathbf{y}(n) \\
 &= \lambda \frac{1}{2} \left[\mathbf{P}(n-1) - \mathbf{k}(n) \mathbf{y}'(n) \mathbf{P}(n-1) \right] \hat{\mathbf{r}}_{dY}(n-1) + d(n) \mathbf{P}(n) \mathbf{y}(n) \\
 &= \mathbf{h}(n-1) - \mathbf{k}(n) \mathbf{y}'(n) \mathbf{h}(n-1) + d(n) \mathbf{k}(n) \\
 &= \mathbf{h}(n-1) + \mathbf{k}(n) \left(d(n) - \mathbf{y}'(n) \mathbf{h}(n-1) \right) \\
 \therefore \mathbf{h}(n) &= \mathbf{h}(n-1) + \mathbf{k}(n) \left(d(n) - \mathbf{y}'(n) \mathbf{h}(n-1) \right)
 \end{aligned}$$

Now, let us see how to get the update of the filter. So \mathbf{h}_n is equal to $\hat{\mathbf{R}}_Y(n)$ inverse into $\hat{\mathbf{r}}_{dY}(n)$. Now this one, we can write as \mathbf{P}_n , because this is our notation \mathbf{P}_n into $\hat{\mathbf{r}}_{dY}(n-1) + d_n$ into \mathbf{y}_n . So that way, we will get this expression is equal to λ times \mathbf{P}_n into $\hat{\mathbf{r}}_{dY}(n-1) + d_n$ into \mathbf{P}_n into \mathbf{y}_n . Now we know the expression for \mathbf{P}_n that expression we will use.

So this expression, we will be using. So λ into $\frac{1}{2}$ by λ times $\mathbf{P}_{n-1} - \mathbf{k}_n$ into \mathbf{y}_n transpose into \mathbf{P}_{n-1} into this $\hat{\mathbf{R}}_{dY}(n-1)$ plus again $d_n \mathbf{y}_n$ into \mathbf{P}_n . So that we can write as d_n into \mathbf{P}_n into \mathbf{y}_n and therefore now, this quantity \mathbf{P}_{n-1} into $\hat{\mathbf{r}}_{dY}(n-1)$ that will be \mathbf{A}_{n-1} , because this is the inverse of the autocorrelation matrix multiplied by the cross correlation function that will be $\mathbf{h}(n-1)$ and this is \mathbf{k}_n and similarly this is \mathbf{P}_{n-1} into $\hat{\mathbf{r}}_{dY}(n-1)$.

That will be again \mathbf{h}_{n-1} and this part \mathbf{P}_n into \mathbf{y}_n that we have established that it is equal to \mathbf{k}_n . So here, we have established that \mathbf{P}_n into \mathbf{y}_n is equal to \mathbf{k}_n . Now we can take \mathbf{k}_n common here from this expression and this expression; therefore, we will get $\mathbf{h}(n-1) + \mathbf{k}_n$ times, this is $d_n - \mathbf{y}'(n) \mathbf{h}(n-1)$. Now this is the desired signal and this is $\mathbf{y}'(n) \mathbf{h}(n-1)$ that is the filtered output, because of the filter $\mathbf{h}(n-1)$ and the difference will be the error.

This error is scaled by gain \mathbf{k}_n and then added to the earlier estimate of \mathbf{h}_n . So that way, \mathbf{h}_n is equal to \mathbf{h}_{n-1} plus \mathbf{k}_n times some prediction error. So this interpretation is helpful. Later on, we will see similar results for (()) (34:46) filter. Therefore, this is like prediction error, because this

you are getting by filtering with previous filter coefficients. So that way this is the predicted output and then this is the desired output.

That difference is the prediction error that is scaled by kn and you are adding to h_{n-1} to get h_n . So that way we have also the recursive relationship for h_n . This is the recursive relationship for h_n . So we have 3 recursion relation now for h_n and for kn and P_n .

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Initialization of P and h

- ❖ We have to initialize $P(-1)$ and $h(-1)$
- ❖ Since P is inverse of an autocorrelation matrix, it is positive definite matrix. Thus a positive definite matrix is chosen for $P(-1)$. For faster convergence,

$$P(-1) = \frac{1}{\delta} I_{M \times M}, \quad \delta \text{ a small positive number}$$

- ❖ $h(-1)$ is initialized at 0.

So these recursive relations are there, but we have to initialize P_n , P matrix and h vector. We have to initialize P of minus 1, because at $n = 0$, we will need P of minus 1 and h of minus 1. Since P is inverse of an autocorrelation matrix, it is positive definite matrix. Thus, a positive definite matrix is chosen for P of minus 1. For faster convergence, we write P of minus 1 is equal to 1 by δ into $I_{M \times M}$, where this is the identity matrix divided by δ .

If we choose δ is small positive number, then 1 by δ will be large. Therefore, this will be a diagonal matrix with large diagonal component. This is required for faster convergence that we will see again later and h of minus 1 is initialized at 0.

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RLS algorithm Steps

Initialization:

$$\mathbf{P}(-1) = \frac{1}{\delta} \mathbf{I}_{M \times M}, \quad \delta \text{ a small positive number}$$

$$\mathbf{h}(-1) = \mathbf{0} \text{ and choose } \lambda$$

Operation:

For $n = 1, 2, \dots$ do

1. Get $d(n), \mathbf{y}(n)$

2. Get $e(n) = d(n) - \mathbf{h}'(n-1)\mathbf{y}(n)$

3. Calculate gain vector $\mathbf{k}(n) = \frac{\mathbf{P}(n-1)\mathbf{y}(n)}{\lambda + \mathbf{y}'(n)\mathbf{P}(n-1)\mathbf{y}(n)}$

4. Update the filter parameters

$$\mathbf{h}(n) = \mathbf{h}(n-1) + \mathbf{k}(n)e(n)$$

5. Update the \mathbf{P} matrix

$$\mathbf{P}(n) = \frac{1}{\lambda} (\mathbf{P}(n-1) - \mathbf{k}(n)\mathbf{y}'(n)\mathbf{P}(n-1))$$

Therefore, the RLS algorithm can be presented in the following steps. First is initialization, that inverse of the autocorrelation matrix that is initialized at 1 by delta times IMM, then h of minus 1 is equal to 0 and we have to choose proper value of lambda, that is the forgetting vector. Now operations are for $n = 0, 1$, etc. What we will do? Get d_n and y_n ; this is the desired signal. This is the input data, then find out the error $e_n = d_n - \mathbf{h}^T(n-1)\mathbf{y}(n)$.

Calculate the gain vector \mathbf{k}_n that is this expression $\mathbf{P}(n-1)\mathbf{y}(n)$ divided by $\lambda + \mathbf{y}_n^T \mathbf{P}(n-1)\mathbf{y}_n$. Then update the filter parameters \mathbf{h}_n is equal to $\mathbf{h}(n-1) + \mathbf{k}_n e_n$ and then update the P matrix by this relation, \mathbf{P}_n is equal to $\frac{1}{\lambda} (\mathbf{P}(n-1) - \mathbf{k}_n \mathbf{y}_n^T \mathbf{P}(n-1))$. So this way this RLS algorithm can be implemented.

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Summary

- ❖ The RLS algorithm minimizes the weighted SSE given by

$$\varepsilon(n) = \sum_{k=0}^n \lambda^{n-k} e^2(k) = \sum_{k=0}^n \lambda^{n-k} (d(k) - \mathbf{y}'(k)\mathbf{h}(n))^2$$

where $e(k)$ is the filtering error at instant k due to $\mathbf{h}(n)$.

- ❖ $0 < \lambda \leq 1$ is the forgetting factor, used to take care of the non-stationarity of the data.

- ❖ After minimization and defining $\hat{\mathbf{R}}_Y(n) = \sum_{k=0}^n \lambda^{n-k} \mathbf{y}(k)\mathbf{y}'(k)$ and

$$\hat{\mathbf{r}}_{rY}(n) = \sum_{k=0}^n \lambda^{n-k} d(k)\mathbf{y}(k),$$

we get the normal equation as

$$\hat{\mathbf{R}}_Y(n)\mathbf{h}(n) = \hat{\mathbf{r}}_{rY}(n)$$

- ❖ The solution to the normal equation is given by

$$\mathbf{h}(n) = \hat{\mathbf{R}}_Y^{-1}(n)\hat{\mathbf{r}}_{rY}(n)$$

Let us summarize the lecture. The RLS algorithm minimizes the weighted sum square error is equal to summation lambda to the power n - k into e square k, k going from 0 to n, which is equal to summation lambda to the power n - k into dk - yk transpose into hn whole square k going from 0 to n, where ek is the filtering error at instant k due to filter at instant and this is very important observation, because you have to use the filter at instant n to filter all the previous data.

Zero less than lambda less than equal to 1 is the forgetting factor used to take care of non-stationarity, for non-stationarity lambda is less than 1 usually 0.99 like that and for stationarity we can take lambda equal to 1. After minimization and defining this quantity Ry hat n is equal to this. This is the weighted sum of this matrix and this is the estimator for the autocorrelation matrix. Similarly, this part rdy hat n is defined by this quantity.

This is the estimator for the cross correlation function. So we get the normal equation Ry hat n into hn is equal to rdy hat n. We get the normal equation as on this autocorrelation matrix into hn is equal to cross correlation vector. The solution to the normal equation, this is the normal equation, is given by hn is equal to because we have to invert this ry hat inverse n into rdy hat n vector. So that way, we can find out hn.

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Summary..

The solution of the normal equation can be written as

$$\mathbf{h}(n) = \left(\lambda \hat{\mathbf{R}}_y(n-1) + \mathbf{y}(n)\mathbf{y}(n)' \right)^{-1} \hat{\mathbf{r}}_x(n)$$

❖ Matrix inversion lemma is applied to get the filter coefficients in three steps:

1. Calculate gain vector $\mathbf{k}(n) = \frac{\mathbf{P}(n-1)\mathbf{y}(n)}{\lambda + \mathbf{y}'(n)\mathbf{P}(n-1)\mathbf{y}(n)}$
2. Update the filter parameters $\mathbf{h}(n) = \mathbf{h}(n-1) + \mathbf{k}(n)e(n)$
3. Update the \mathbf{P} matrix $\mathbf{P}(n) = \frac{1}{\lambda} \left(\mathbf{P}(n-1) - \mathbf{k}(n)\mathbf{y}'(n)\mathbf{P}(n-1) \right)$

where $\mathbf{P}(n) = \hat{\mathbf{R}}_y^{-1}(n)$

This solution of the normal equation can be written as $\mathbf{h}(n)$ because using the recursive relation for $\hat{\mathbf{R}}_y(n)$, $\mathbf{h}(n)$ can be written as $\lambda \hat{\mathbf{R}}_y(n-1) + \mathbf{y}(n)\mathbf{y}(n)'$ whole inverse into $\hat{\mathbf{r}}_x(n)$. So that way we can write $\mathbf{h}(n)$ and now we can do the matrix inversion lemma is applied to get the filter coefficient in three steps. So what are these three steps: calculate the gain vector $\mathbf{k}(n)$ given by $\mathbf{P}(n-1)$ into $\mathbf{y}(n)$ divided by $\lambda + \mathbf{y}(n)'$ into $\mathbf{P}(n-1)$ into $\mathbf{y}(n)$.

This is the gain vector. Update the filter parameters $\mathbf{h}(n) = \mathbf{h}(n-1) + \mathbf{k}(n)e(n)$, where $e(n)$ is the difference between $d(n)$ and the filtered output. That we have shown earlier. Then we have to update the \mathbf{P} matrix by $\mathbf{P}(n) = \frac{1}{\lambda} \left(\mathbf{P}(n-1) - \mathbf{k}(n)\mathbf{y}(n)'\mathbf{P}(n-1) \right)$ and in these steps $\mathbf{P}(n)$ is equal to inverse of the estimated autocorrelation matrix. Thus, we have derived the RLS algorithm. In the next lecture, we will discuss some important properties of RLS adaptive filters and compare those properties with those of LMS filters. Thank you.